

Brownian Motion

1.

1. Conditioning of gaussian vectors

let $Y, X_1, \dots, X_n$ be a gaussian vector. Compute $\hat{\alpha}_1, \dots, \hat{\alpha}_n$ such that

$Y - \sum_{j=1}^n \hat{\alpha}_j X_j$ is uncorrelated with $X_j, j=1, \dots, n$.

the condition is

$$\begin{aligned} \text{Cov} \left(Y - \sum_{j=1}^n \hat{\alpha}_j X_j, X_i \right) &= \text{Cov}(Y, X_i) - \\ &\sum_{j=1}^n \hat{\alpha}_j \text{Cov}(X_j, X_i) = 0, \quad i=1, \dots, n. \end{aligned}$$

$$\text{Th. 4} \quad \mathbb{E}(Y | X_1, \dots, X_n) = \mathbb{E}(Y) + \sum_{j=1}^n \hat{\alpha}_j (X_j - \mathbb{E}(X_j))$$

Proof. $\mathbb{E}(Y | X_1, \dots, X_n) =$

$$\begin{aligned} &\mathbb{E} \left(Y - \sum_{j=1}^n \hat{\alpha}_j X_j + \sum_{j=1}^n \hat{\alpha}_j X_j \mid X_1, \dots, X_n \right) = \\ &\mathbb{E} \left(Y - \sum_{j=1}^n \hat{\alpha}_j X_j \right) + \sum_{j=1}^n \hat{\alpha}_j X_j \quad \square \end{aligned}$$

We have used

1) Uncorrelated gaussian vectors are independent

2) If X and Y are independent, then

$$\mathbb{E}(X | Y) = \mathbb{E}(X)$$

Lemma | let X independent of Y ($X \perp Y$)

then

$$\mathbb{E}(\phi(X, Y) | X) = \int \phi(X, y) F_Y(dy)$$

Proof. F_Y is the distribution of the random

variable Y . Define $\hat{\phi}(x) = \int \phi(x, y) F_Y(dy)$

$= \mathbb{E}(\phi(x, Y))$. The random variable

$\hat{\phi}(X) = \int \phi(X, y) F_Y(dy)$ is equal to

$\mathbb{E}(\phi(x, Y) | X) \stackrel{\text{i.p.t.}}{=} \int \phi(x, y) \psi(y) F_Y(dy)$ for all function $\psi(x)$ 2.

$$\mathbb{E}(\phi(x, Y) \psi(x)) = \mathbb{E}(\hat{\phi}(x) \psi(x))$$

Because of $X \perp\!\!\!\perp Y$, the lhs is

$$\int \int \phi(x, y) \psi(x) F_X(dx) F_Y(dy) =$$

$$\int \left(\int \phi(x, y) F_Y(dy) \right) \psi(x) F_X(dx) = \int \hat{\phi}(x) \psi(x) F_X(dx) \quad \square$$

Th. 2 Under the same conditions as the th. 1

$$\mathbb{E}(\phi(Y) | X_1, \dots, X_n) =$$

$$\int \phi\left(z + \sum_{j=1}^n \hat{\alpha}_j X_j\right) F_{Y - \sum_{j=1}^n \hat{\alpha}_j X_j}(dz)$$

Proof.

$$\begin{aligned} \mathbb{E}(\phi(Y) | X_1, \dots, X_n) &= \\ \mathbb{E}\left(\phi\left(Y - \sum_{j=1}^n \hat{\alpha}_j X_j + \sum_{j=1}^n \hat{\alpha}_j X_j\right) | X_1, \dots, X_n\right) &= \\ \int \phi\left(z + \sum_{j=1}^n \hat{\alpha}_j X_j\right) F_{Y - \sum_{j=1}^n \hat{\alpha}_j X_j}(dz) &\quad \square \end{aligned}$$

Ex 1. Assume $\Gamma_X = [\text{Cov}(X_i, X_j)]_{i,j=1,\dots,n}$ non-singular.
Compute the conditional distribution of Y given $X_1, \dots, X_n$

Ex 2. Same with Γ_X singular

2. Brownian motion

DI

(Shreve p. 94) let $(\Omega, \mathcal{F}, P)$ be a probability space and W a continuous stochastic process with continuous trajectories. W is a BM (or a Wiener process WP) if

1) $W_0 = 0$

2) $0 = t_0 < t_1 < \dots < t_n$ the increments

$$W_{t_1}, W_{t_2} - W_{t_1}, \dots, W_{t_n} - W_{t_{n-1}}$$

are independent ~~and~~ ³ and normally distributed with

$$\begin{cases} \mathbb{E}(W_t - W_s) = 0 \\ \text{Var}(W_t - W_s) = t - s \end{cases} \quad s < t$$

If we define $\mathcal{F}_t = \sigma(W_s, s \leq t)$, then the process $(W_t)_{t \geq 0}$ is adapted to $(\mathcal{F}_t)_{t \geq 0}$.

th (Wiener) $\Omega = C(\mathbb{R}_+)$, $W_t(\omega) = \omega(t)$,
 there exist P (the Wiener probability measure)
 such that W is a BM.

Ex 2. Compute the joint density of $W_{t_1}, \dots, W_{t_n}$

A more general definition is the following

D2. let $(\Omega, \mathcal{F}, (\mathcal{F}_t)_{t \geq 0}, P)$ a filtered probability space. A process continuous process W is a BM if for all $s < t$

$$W_t - W_s \sim N(0, t-s) \text{ and } W_t - W_s \notin \mathcal{F}_s. \quad 4/$$

Ex 3. Prove the implication i.e. compare D1 with D2.

Properties

1. A BM is a martingale

$$\begin{aligned} \mathbb{E}(W_t | \mathcal{F}_s) &= \mathbb{E}(W_t - W_s + W_s | \mathcal{F}_s) \\ &= \mathbb{E}(W_t - W_s) + W_s = W_s \end{aligned}$$

2. A BM is a Markov process

$$\begin{aligned} \mathbb{E}(\varphi(W_t) | \mathcal{F}_s) &= \mathbb{E}(\varphi(W_t - W_s + W_s) | \mathcal{F}_s) \\ &= \int \varphi(z + W_s) \frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{1}{2(t-s)} z^2} dz \\ &= \int \varphi(y) \frac{1}{\sqrt{2\pi(t-s)}} e^{-\frac{1}{2(t-s)} (y - W_s)^2} dy \quad y = z + W_s \end{aligned}$$

that is the conditional distribution of W_t given $\mathcal{F}_s$ is $N(W_s, t-s)$.

3. Piece-wise linear approximation

4. Let $f: [0, T] \rightarrow \mathbb{R}$ be continuous. The n -th piece-wise linear approximation of f is a continuous function $f_n: [0, T] \rightarrow \mathbb{R}$ such that

$$1) f_n\left(\frac{k}{n}T\right) = f\left(\frac{k}{n}T\right) \quad k=0, 1, \dots, n$$

$$2) f_n(t) = f\left(\frac{k}{n}T\right) + \left(f\left(\frac{k+1}{n}T\right) - f\left(\frac{k}{n}T\right)\right) \frac{t - \frac{k}{n}T}{\frac{k+1}{n}T - \frac{k}{n}T} \quad t \in \left[\frac{k}{n}T, \frac{k+1}{n}T\right] \quad (m: 2 \rightarrow 1)$$

As f is continuous on a bounded interval the modulus of continuity

$$\delta(f, \epsilon) = \sup \left\{ |f(s) - f(t)| : |t - s| \leq \epsilon \right\}$$

is finite and

$$\lim_{\epsilon \rightarrow 0} \delta(f, \epsilon) = 0.$$

We have

$$\begin{aligned} \sup_t |f(t) - f_n(t)| &= \sup_k \sup_{t \in [\frac{k}{n}T, \frac{k+1}{n}T]} |f(t) - f_n(t)| \\ &= \sup_k \sup_{t \in [\frac{k}{n}T, \frac{k+1}{n}T]} \left| \left(f(t) - f\left(\frac{k}{n}T\right)\right) + \left(f\left(\frac{k+1}{n}T\right) - f\left(\frac{k}{n}T\right)\right) \frac{t - \frac{k}{n}T}{\frac{k+1}{n}T - \frac{k}{n}T} \right| \\ &\leq \delta\left(f, \frac{T}{n}\right) + \delta\left(f, \frac{T}{n}\right) \frac{T}{n} = \delta\left(f, \frac{T}{n}\right) \left(1 + \frac{T}{n}\right) \rightarrow 0 \\ &\quad \text{as } n \rightarrow +\infty. \end{aligned}$$

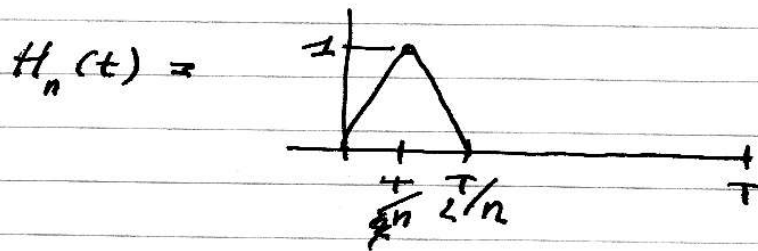
B. Define $W_t^{(n)}(w)$ to be the n -th piece-wise linear approximation of the continuous trajectory $t \mapsto W_t(w)$, $t \in [0, T]$. Then

$$\sup_{t \in [0, T]} |W_t(\omega) - W_t^{(n)}| = 0 \text{ for all } \omega$$

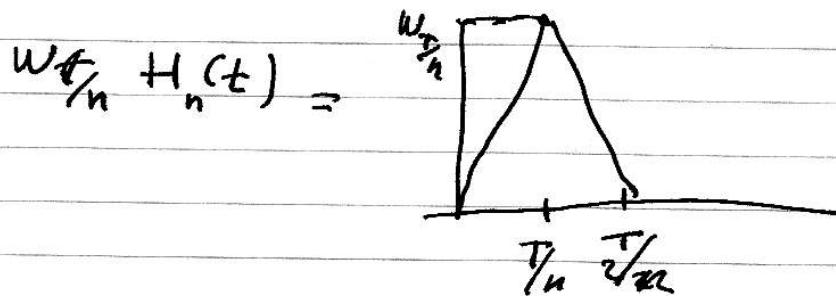
such that $t \mapsto W_t(\omega)$ is continuous.

Note that $W_t^{(n)}$ is a Gaussian process.

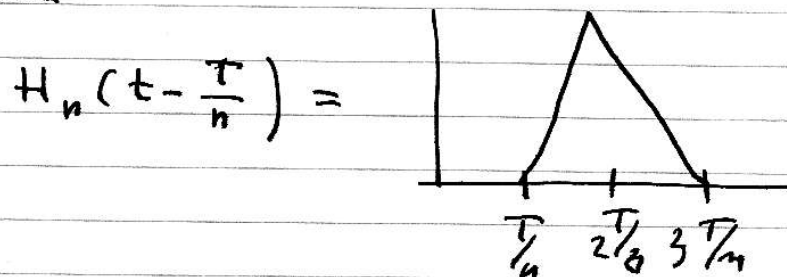
c. Consider the function ("wavelet")



then

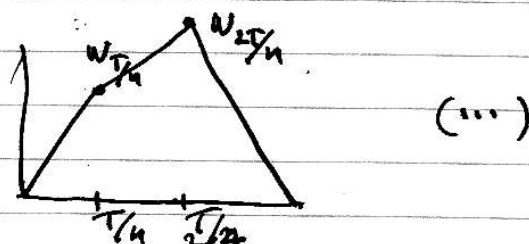


We have



and

$$W_{T/n} H_n(t) + W_{2T/n} H_n(t - \frac{T}{n}) + \dots$$



In general

$$W_t^{(n)} = \sum_{k=1}^n \frac{W_{k/n} - W_{(k-1)/n}}{\sqrt{T/n}} H_n(t - \frac{k-1}{n}T) \quad t \in [0, T] \quad \text{✓}$$

In term of the increments, we have

$$\frac{W_{k/n} - W_{(k-1)/n}}{\sqrt{T/n}} = \sqrt{\frac{T}{n}} Z_k \quad \text{with}$$

$Z_k \quad k=1, \dots, n \quad \text{i.i.d. } N(0,1)$. By adding all increments,

$$\frac{W_{k/n} - W_{(k-1)/n}}{\sqrt{T/n}} = \sqrt{\frac{T}{n}} (Z_1 + \dots + Z_k)$$

$$W_t^{(n)} = \sum_{k=1}^n \sqrt{\frac{T}{n}} \sum_{h=1}^k Z_h H_n(t - \frac{k-1}{n}T)$$

$$\sqrt{\frac{T}{n}} \sum_{h=1}^n Z_h \underbrace{\sum_{k=h}^n H_n(t - \frac{k-1}{n}T)}_{\text{Ex 4 ?}}$$

Ex 4 ?

Ex 5. Compute $\text{Cov}(W_s^{(n)}, W_t^{(n)})$ $s \leq t$

D. Recursive simulation of $w^{(1)}, w^{(2)}, \dots, w^{(n)}, \dots$

Assume $T=1$. The sequence $Z_1, Z_2, \dots$ is i.i.d $N(0,1)$ and (potentially) infinite.

Then $Z_1 \sim W_1$ and

$$W_t^{(1)} \sim t Z_1$$

4. Quadratic variation

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Consider a partition of $[0, t]$, $t_k = \frac{k}{n}t$, $k=0, 1, \dots, n$.

$$\sum_{k=1}^n (W_{t_k} - W_{t_{k-1}})^2 = Q_t^{(n)}. \quad \text{Each}$$

$$W_{t_k} - W_{t_{k-1}} \sim N(0, \frac{t}{n}), \quad \text{i.e.}$$

$$W_{t_k} - W_{t_{k-1}} = \sqrt{\frac{t}{n}} Z_k$$

$$Z_1, \dots, Z_n \text{ i.i.d. } N(0, 1)$$

$$Q_t^{(n)} = \sum_{k=1}^n \left(\sqrt{\frac{t}{n}} Z_k \right)^2 = \frac{t}{n} \sum_{k=1}^n Z_k^2$$

By a version of the law of large numbers

$$Q_t^{(n)} \rightarrow t \quad \mathbb{E}(Z_1^2) = 1$$

Th. 1) W_t^2 is a sub-martingale, i.e.

$$\mathbb{E}(W_t^2 | \mathcal{F}_s) \geq W_s^2$$

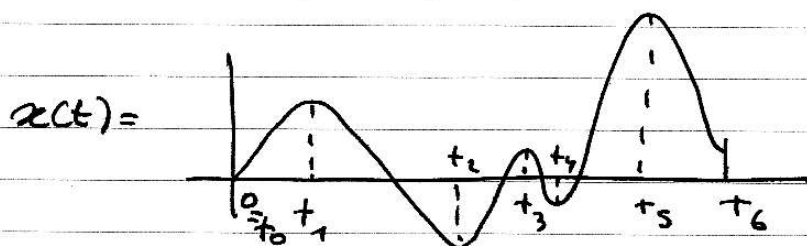
2) $W_t^2 - t$ is a martingale

3) $e^{2W_t - \frac{a^2}{2}t}$ is a Martingale

Proof Ex. 7.

Variation of a function

the variation of the trajectory



is

$$|x(t_1) - x(t_0)| + |x(t_2) - x(t_1)| + |x(t_3) - x(t_2)| \\ + |x(t_4) - x(t_3)| + |x(t_5) - x(t_4)| + |x(t_6) - x(t_5)|$$

More generally, we check the existence of the limit along the partitions

$\pi = t_0=0, t_1, t_2, \dots, t_n=t$
of the total sum of absolute increments

$$\lim_{n \rightarrow \infty} \sum_{i=1}^n |x(t_i) - x(t_{i-1})|$$

Assume $f \in C^1$. By Lagrange (!) MVT

~~$$x(t_i) - x(t_{i-1}) = x'(s_i)(t_i - t_{i-1})$$~~

$$x(t_i) - x(t_{i-1}) = x'(s_i)(t_i - t_{i-1})$$

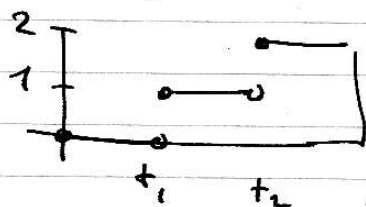
and

$$\sum_{i=1}^n |x'(s_i)|(t_i - t_{i-1}) \text{ is convergent}$$

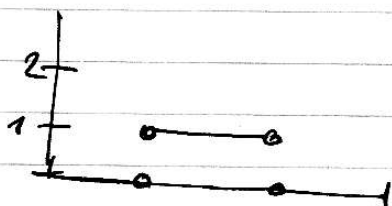
to

$$\int_0^t |x'(s)| ds.$$

Finite variation does not require continuity. 5/
 A step function (trajectory) has finite variation, e.g.



has variation = 2, as it is non-decreasing, and eq. has variation 2.



this is of interest for processes with jumps.

It is more convenient to free the running time t from the partition, i.e. we take

$$\pi = 0, t_1 < t_2 < \dots < t_n < \dots \quad t_n \rightarrow +\infty$$

and compute the variation at each time t as

$$\lim_{\pi} \sum_i |x(t_i) - x(t_{i-1})|$$

as

$$\text{for } t \leq t_{i-1} \quad \text{we have} \quad x(t) - x(t) = 0$$

$$\text{" } t_{i-1} < t < t_i \quad \text{" } x(t) - x(t_{i-1})$$

$$\text{" } t_i \leq t \quad \text{" } x(t_i) - x(t_{i-1})$$

Basic computation: Assume x finite variation and continuous. Fix a partition. By telescoping and Taylor formula:

$$\begin{aligned}
 x(t)^2 - x(0)^2 &= \sum_{i=1}^{\infty} x(t_i)^2 - x(t_{i-1})^2 \quad 6/ \\
 &= \sum_{i=1}^{\infty} 2 x(t_{i-1}) (x(t_i) - x(t_{i-1})) \\
 &\quad + \sum_{i=1}^{\infty} |x(t_i) - x(t_{i-1})|^2.
 \end{aligned}$$

The last term is bounded by

$$\sup_i |x(t_i) - x(t_{i-1})| \times \sum_{i=1}^{\infty} |x(t_i) - x(t_{i-1})|$$

where the first factor goes to zero (continuity)
and the second factor is finite (finite variation)

It follows the existence of the limit

$$\begin{aligned}
 \lim_{n \rightarrow \infty} 2 \sum_{i=1}^n x(t_{i-1}) (x(t_i) - x(t_{i-1})) \\
 \equiv 2 \int_0^t x(t) dx(t)
 \end{aligned}$$

and the formula

$$x^2(t) - x(0)^2 = 2 \int_0^t x(t) dx(t)$$

Ex 3. Compute $\varphi(x(t))$ where $\varphi \in C^2$ with bounded derivatives

Quadratic Variation

W is a Wiener Process (Brownian motion) on $(\Omega, \mathcal{F}, \mathbb{P}, (\mathcal{F}_t)_{t \geq 0})$. Fix a partition

$$0 = t_0 < t_1 < \dots < t_n = t$$

and compute (telescoping and Taylor expansion)

$$\begin{aligned} W_t^2 - W_0^2 &= \sum_{i=1}^n W_{t_{i+1}}^2 - W_{t_i}^2 \\ &= 2 \sum_{i=1}^n W_{t_i} (W_{t_{i+1}} - W_{t_i}) + \sum_{i=1}^n (W_{t_{i+1}} - W_{t_i})^2 \end{aligned}$$

Both LHS and RHS are stochastic processes. In order to see the dependence on t in the RHS we change the notation as follows

$$0 = t_0 < t_1 < t_2 < \dots < t_n < t_{n+1} < \dots$$

with $t_n \rightarrow \infty$ as $n \rightarrow \infty$ and write

$$\begin{aligned} W_t^2 - W_0^2 &= \sum_i W_{t \wedge t_{i+1}}^2 - W_{t \wedge t_i}^2 \\ &= \sum_i 2 W_{t \wedge t_i} (W_{t \wedge t_{i+1}} - W_{t \wedge t_i}) + \sum_i (W_{t \wedge t_{i+1}} - W_{t \wedge t_i})^2 \end{aligned}$$

To check the formulae, consider $t \in [t_i, t_{i+1}]$.

The continuous process

$$M_t^n = \sum_{i=1}^n 2 W_{t \wedge t_i} (W_{t \wedge t_{i+1}} - W_{t \wedge t_i})$$

is a MARTINGALE. In fact consider one term of the sum and check the conditioning in each case:

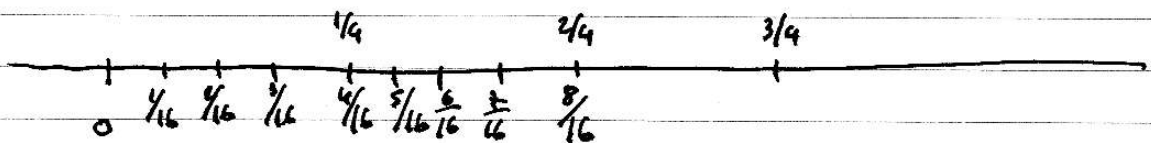
$$2W_{t \wedge t_i} (W_{t \wedge t_{i+1}} - W_{t \wedge t_i}) = \begin{cases} 0 & t \leq t_i \\ 2W_{t_i} (W_t - W_{t_i}) & t_i < t \leq t_{i+1} \\ 2W_{t_{i+1}} (W_{t_{i+1}} - W_{t_i}) & t_{i+1} < t \end{cases}$$

let us compare two approximations, one of step 2^{-m} and the other with step 2^{-n} , $m < n$.

$$M_t^{(m)} = \sum_k 2W_{t \wedge \frac{k}{2^m}} (W_{t \wedge \frac{k+1}{2^m}} - W_{t \wedge \frac{k}{2^m}})$$

$$M_t^{(n)} = \sum_k 2W_{t \wedge \frac{k}{2^n}} (W_{t \wedge \frac{k+1}{2^n}} - W_{t \wedge \frac{k}{2^n}})$$

$$M_t^{(n)} - M_t^{(m)} = \sum_{k=1}^{\infty} 2 \left(W_{t \wedge \frac{k}{2^n}} - W_{t \wedge \frac{k}{2^m}} \right) \left(W_{t \wedge \frac{k+1}{2^n}} - W_{t \wedge \frac{k}{2^n}} \right)$$



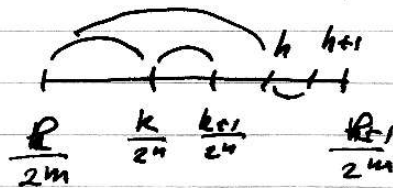
let us compute $E (M_t^{(n)} - M_t^{(m)})^2 =$

$$\sum_{k,k'=1}^{\infty} 4 E \left[\left(W_{t \wedge \frac{k}{2^n}} - W_{t \wedge \frac{k'}{2^n}} \right) \left(W_{t \wedge \frac{k+1}{2^n}} - W_{t \wedge \frac{k}{2^n}} \right) \right] \\ E \left(W_{t \wedge \frac{k}{2^n}} - W_{t \wedge \frac{k'}{2^n}} \right) \left(W_{t \wedge \frac{k+1}{2^n}} - W_{t \wedge \frac{k}{2^n}} \right)$$

Case ~~h < k~~ $h = k$

$$= \sum_{k=1}^{\infty} 4 \mathbb{E} \left[\left(W_{t \wedge \frac{k}{2^n}} - W_{t \wedge \frac{k+1}{2^n}} \right)^2 \left(V_{t \wedge \frac{k+1}{2^n}} - V_{t \wedge \frac{k}{2^n}} \right)^2 \right]$$

Case $h \neq k$:



$$= 4 \sum_{k=1}^{\infty} \mathbb{E} \left(W_{t \wedge \frac{k}{2^n}} - W_{t \wedge \frac{k+1}{2^n}} \right) * \mathbb{E} \left(t \wedge \frac{k+1}{2^n} - t \wedge \frac{k}{2^n} \right)$$

$$\left(t \wedge \frac{k}{2^n} - t \wedge \frac{k}{2^n} \right)$$

$$\leq 4 t 2^{-m} \quad \text{Ex 4. Check the proof in Shreve p.103}$$

Conclusion

1. The martingale $M^{(n)} - M^{(m)}$ is such that $\mathbb{E} \left((M_t^{(n)} - M_t^{(m)})^2 \right) \rightarrow 0$ as $m, n \rightarrow \infty$

2. The Doob maximal lemma implies

$$\mathbb{E} \left(\sup_{0 \leq t} |M_t^{(n)} - M_t^{(m)}|^2 \right) \rightarrow 0 \text{ as } m, n \rightarrow \infty$$

3. The limit $M^{(n)} \rightarrow 2 \int_0^t W_s dW_s$ is a continuous, adapted MARTINGALE.

4. The processes $\sum_i (W_{t \wedge t_i} - W_{t \wedge t_{i-1}})^2$ converge in the sense 2. to the quadratic variation t which is a compensator of W_t^2 .