COLLEGIO CARLO ALBERTO STOCHASTIC PROCESSES 2014

3. INFINITELY DIVISIBLE DISTRIBUTIONS

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The following lecture notes are based on Sasvári [2] and Sato [3].

1. **Definition** (Convolution). Let X, Y be independent random variables with values in \mathbb{R}^n and distributions μ_X , μ_Y , respectively. The convolution $\mu_X * \mu_Y$ is the distribution of X + Y, that is for all bounded $f : \mathbb{R}^n \to \mathbb{R}$

$$\int f(z) \ (\mu_X * \mu_Y)(dz) = \operatorname{E}\left(f(X+Y)\right) = \iint f(x+y) \ \mu_X(dx)\mu_Y(dy).$$

- 2. Proposition (Convolution).
 - (1) If μ_1 has density p_1 with respect to the Lebesgue measure, then $\mu_1 * \mu_2$ has density $p_1 * \mu_2$ given by

$$p_1 * \mu_2(z) = \int p_1(z-y) \ \mu_2(dy).$$

(2) If both have densities p_1 , p_2 , respectively, with respect to the Lebesgue measure, then $\mu_1 * \mu_2$ has density $p_1 * p_2$ given by

$$(p_1 * p_2)(z) = \int p_1(z-y)p_2(y) \, dy = \int p_1(x)p_2(z-x) \, dx.$$

(3) If the measurable function $f : \mathbb{R}^n \to \mathbb{R}$, and μ is a measure, define $f * \mu(z) = \int f(z-x) \ \mu(dx)$ if the integral exists a.e, i.e. $\int |f(z-x)| \ \mu(dx) < \infty$ a.s in z. If the measure μ is finite and $f \in L^a(dx)$, $1 \leq a \leq \infty$, then then $f * \mu$ exists and $\|f * \mu\|_a \leq \|f\|_a$.

Exercise.

(1) For each bounded $f : \mathbb{R}^n \to \mathbb{R}$

$$\int f(z) \ (\mu_1 * \mu_2)(dz) = \iint f(x+y) \ p_1(x) \ dx \ \mu_2(dy) = \\ \int \left(\int f(x+y) \ p_1(x) \ dy \right) \ \mu_2(dy) = \int \left(\int f(z) \ p_1(z-x) \ dz \right) \ \mu_2(dy) = \\ \int f(z) \left(\int \ p_1(z-y) \ \mu_2(dy) \right) \ dz.$$

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(2) For each bounded $f : \mathbb{R}^n \to \mathbb{R}$

$$\int f(z) \ (\mu_1 * \mu_2)(dz) = \iint f(x+y) \ p_1(x)dx \ p_2(y)dy = \\ \int \left(\int f(x+y) \ p_2(y)dy \right) \ p_1(x)dx = \int \left(\int f(z) \ p_2(z-x)dz \right) \ p_1(x)dx = \\ \int f(z) \left(\int p_1(x)p_2(z-x)dx \right) \ dz.$$
(3) From $\iint |f(z-x)|^a \ dx\mu(dz) = \int |f(y)|^a \ dy.$

3. Exercise.

- (1) Compute $\chi * f$ for $\chi = \delta_a$ and $f \in C_b$.
- (2) Compute $\chi * f$ for $\chi = (\delta_a \delta_b)^{*n}$, n = 1, 2, 3 and $f \in C_b$. If f is a polynomial of degree n 1, then $(\delta_a \delta_b)^{*n} * f = 0$.

- (3) Compute $\chi * f$ for $\chi(dx) = A^{-1}(0 \le x \le A)dx$. If χ is a complex measure, define $\widetilde{\chi}(B) = \overline{\chi(-B)}$. Compute $\chi * \widetilde{\chi}$ and $\chi * \widetilde{\chi} * f$, $f \in \mathbb{C}_b$.
- (4) Let $mu_{\sigma} = N(0, \sigma^2 I)$, and $f \in C_b$. Then $\lim_{\sigma \to 0} f * \mu_{\sigma}(z) = f(z)$.

Solution.

- (1) $\delta_a * f(y) = \int f(y-x) \, \delta_a(dx) = f(y-a)$ is the translation of f.
- (2) $(\delta_a \delta_b) * f(y) = f(y-a) f(y-b)$. If f is constant, then f(y-a) f(y-b) = 0. $(\delta_a - \delta_b)^{*2} * f(y) = f(y-2a) - 2f(y-a-b) + f(y-2b)$. If $f = u^t y + v$ is a polynomial of degree 1, then $f(y-a) - f(y-b) = u^t(y-a) - u^t(y-b) = u^t(b-a)$ is constant. In general, if f is a polynomial of degree n, then $(\delta_a - \delta_b) * f(y) = f(y-a) - f(y-b) = \sum_{k=1}^{x} \frac{1}{k!} f^{(k)}(y-b)(b-a)^k$ is a polynomial of degree at most n-1, e.g. $(y-a)^2 - (y-b)^2 = a(y-b)(b-a) + (a-b)^2$.
- (3) Continuity and dominated convergence:

$$f * \mu_{\sigma}(y) = \int f(y-x)(2\pi)^{-n/2} \sigma^{-n} e^{-\|x\|^2/2\sigma^2} dy = \int f(y-\sigma z)(2\pi)^{-n/2} e^{-\|x\|^2/2} dz \to \int f(y)(2\pi)^{-n/2} e^{-\|x\|^2/2} dz = f(y).$$

- 4. **Definition** (Characteristic function, inverse Fourier tranform).
 - (1) Let X be a random variable in \mathbb{R}^n , with distribution μ_X . The characteristic function of X is the function $\phi_X : \mathbb{R}^n \to \mathbb{C}$ defined by

$$\phi_X(t) = \mathcal{E}\left(e^{i\langle t,X\rangle}\right) = \int e^{i\langle t,x\rangle} \mu(dx), \quad i = \sqrt{-1}.$$

(2) The function $\check{\mu}_X(t) = \int e^{i\langle t,x \rangle} \mu_{(}dx)$ is also called inverse Fourier transform of μ . If μ has a density f with respect to the Lebesgue measure, then the inverse Fourier transform of f is $\check{f}(t) = \int e^{i\langle t,x \rangle} f(x) \, dx = \check{\mu}(t)$. If $f \in L^1(dx)$ then the Fourier transform $\hat{f}(t) = \int e^{-i\langle t,x \rangle} f(x) \, dx$ is defined for all t and $\|\hat{f}\|_1 \leq \|f\|_1$.

5. **Definition** (Positive definite). A function $f: \mathbb{R}^n \to \mathbb{C}$ is positive definite if for all $m \in \mathbb{N}, x_1, \ldots, x_m \in \mathbb{R}^n, c_1, \ldots, c_m \in \mathbb{C}$ it holds $\sum_{i,j=1}^n c_i \overline{c_j} f(x_i - x_j) \ge 0$. In other words, for all $m \in \mathbb{N}, x_1, \ldots, x_m \in \mathbb{R}^n$ the matrix $A = [f(x_i - x_j)]_{i,j=1}^m$ is positive definite, that is for all $u \in \mathbb{C}^m$, it holds $u^*Au \ge 0, u^* = \overline{u^t}$.

- 6. Definition (Hermitian). A function $f : \mathbb{R}^n \to \mathbb{C}$ is Hermitian if $f(-t) = \overline{f(t)}, t \in \mathbb{R}^n$.
- 7. Proposition (Properties of positive definite functions).
 - (1) If f is positive definite, then f is Hermitian and $f(0) \ge 0$.
 - (2) For a complex measure μ , define $\tilde{\mu}$ by $\int f(x) \tilde{\mu}(dx) = \overline{\int f(-x) \mu(dx)}$. Analogously, for each $g: \mathbb{R}^n \to \mathbb{C}$, define $\tilde{g}(t) = \overline{g(-t)}$. If f is continuous and positive definite, then $\mu * \widetilde{\mu} * f$ is positive definite. Similarly for $g * \widetilde{g} * f(t)$.
 - (3) If $g \in L^2(\mathbb{R}^n, \lambda; \mathbb{C})$, that is $\int |g(x)|^2 dx < \infty$, then $g * \widetilde{g}$ exists and is positive definite.
 - (4) Let $X_t, t \in \mathbb{R}^n$ be a family of random variables such that $\text{Cov}(X_t, X_s) = \rho(t-s)$. Then ρ is positive definite.
 - (5) If f is positive definite, then $x \mapsto e^{i\langle t,x \rangle} f(x)$ is positive definite.
 - (6) If f is positive definite and integrable, the $\int f(x) dx \ge 0$.

Exercise.

- (1) Take m = 1 $c_1 = 1$, $x_1 = 0$: $c_1\overline{c_1}f(0-0) = f(0) \ge 0$. Take m = 2, $x_1 = 0$, $x_2 = x$: $|c_1|^2 f(0) + c_1 \overline{c_2} f(-x) + c_2 \overline{c_1} f(x) + |c_2| f(0) \ge 0$. In particular, with $c_1 = c_2 = 1$, we have $f(-x) + f(x) \in \mathbb{R}$, and, with $c_1 = 1$, $c_2 = i$, we have $i(f(x) - f(-x)) \in \mathbb{R}$. It follows $f(x) = [(f(-x) + f(x)) - i \cdot i(f(x) - f(-x))]/2$ and $f(-x) = [(f(-x) + f(x)) + i \cdot i(f(x) - f(-x))]/2.$ (2) It is enough to consider $\mu = \sum_{i=1}^{m} c_i \delta_{x_j}, \ \widetilde{\mu} = \sum_{j=1}^{m} \overline{c_j} \delta_{-x_j}.$ In such a case

$$\mu * \widetilde{\mu} * f(z) = \iint f(z - x - y) \ \mu(dx)\widetilde{\mu}(dy) = \sum_{i,j=1}^m c_i \overline{c_j} f(z - x_i + x_j).$$

We check the positive definiteness with

$$\sum_{h,k=1}^{M} d_h \overline{d_k} \mu * \widetilde{\mu} * f(z) = \iint f(z - x - y) \ \mu(dx) \widetilde{\mu}(dy) = \sum_{h,k=1}^{M} d_h \overline{d_k} \sum_{i,j=1}^{m} c_i \overline{c_j} f(z_h - z_k - x_i + x_j) = \sum_{h,k=1}^{M} \sum_{i,j=1}^{m} d_h \overline{d_k} c_i \overline{c_j} f((z_h - x_i) - (z_k - x_j)) = \sum_{(h,i),(k,j)} (c_i d_h) \overline{(c_j d_k)} (f(y_{h,i} - y_{k,j}) \ge 0.$$

(3) The existence of $g * \widetilde{g}(y) = \int g(y-x)\overline{g(-x)} \, dx$ follows from

$$|g * \widetilde{g}(y)| \leq \sqrt{\int |g(y-x)|^2} dx \sqrt{\int \left|\overline{g(-x)}\right|^2} dx = ||g||_2^2.$$

The positive definiteness is

$$\sum_{i,j=1}^{m} c_i \overline{c_j} \int g(y_i - y_j - x) \overline{g(-x)} \, dx = \int \sum_{i,j=1}^{m} c_i \overline{c_j} g(y_i - y_j - x) \overline{g(-x)} \, dx = \int \left| \sum_{i,j=1}^{m} c_i \overline{c_j} g(y_i - x) \overline{g(y_j - x)} \, dx \right| = \int \left| \sum_{i=1}^{m} c_i g(y_i - x) \right| \, dx \ge 0$$

- (4) From the definition.
- (5) From the definition.

(6) Let g_n be a sequence of triangular functions such that $f(z)g_n(z) \to f(z)$. Write $g_n = h_n * \widetilde{h_n}$, h_n being uniform, and compute $\int f(x)\tau_n(x) dx$ as the value at 0 of a positive convolution. I.e. (n=1) let $h_n(x) = n^{-1/2}(0 \le x \le n)$ and define $g_n = h_n * \widetilde{h_n}$. Then $g_n(0) = 1$, $0 \le g_n(x) \le 1$ and $g_n(x) \to 1$, $n \to \infty$. From $\int f(x)g_n(x) dx = f * g_n * \widetilde{g_n}(0) \ge 0$ we obtain the result.

8. Proposition.

- (1) The characteristic function is uniformly continuous and $\check{\mu}(0) = 1$.
- (2) If X is a random variable in \mathbb{R}^n with characteristic function ϕ_X , for each $A \in \mathbb{R}^{m \times n}$ and $a \in \mathbb{R}^m$, the random variable Y = a + AX has characteristic function $\phi_Y(s) = e^{i\langle a,s \rangle} \phi_X(A^T s)$.
- (3) If X_1 and X_2 are independent random variables with values in \mathbb{R}^{n_1} and \mathbb{R}^{n_2} , and characteristic functions ϕ_1 , ϕ_2 , respectively, then $X = (X_1, X_2)$ has characteristic function $\phi_X(t_1, t_2) = \phi_1(t_1)\phi_2(t_2)$.
- (4) The characteristic function is Hermitian.
- (5) The characteristic function is positive definite.

Exercise.

(1) We have
$$|e^{i\langle t+h,x\rangle} - e^{i\langle t,x\rangle}| = |e^{i\langle t,x\rangle}(e^{i\langle h,x\rangle} - 1)| = |e^{i\langle h,x\rangle} - 1| \leq 2$$
, and $\lim_{h\to 0} O(h) = \int |e^{i\langle h,x\rangle} - 1| \ \mu(dx) = 0$ by dominated convergence.

- (2) From $\langle s, a + Ax \rangle = \langle s, a \rangle + \langle A^T s, x \rangle$.
- (3) From independence and $\langle (t_1, t_2), (x_1, x_2) \rangle_{n_1+n_2} = \langle t_1, x_1 \rangle_{n_1} + \langle t_2, x_2 \rangle_{n_2}$. (4) $\check{\mu}(-t) = \int e^{i\langle -t, x \rangle} \mu(dx) = \int \overline{e^{i\langle t, x \rangle}} \mu(dx) = \overline{\check{\mu}(t)}$. (5)

$$\sum_{i,j=1}^{n} c_i \overline{c_j} \check{\mu}(t_1 - t_j) = \int \sum_{i,j=1}^{n} c_i \overline{c_j} e^{i\langle t_i - t_j, x \rangle} \, \mu(dx) = \int \left| \sum_{i,j=1}^{n} c_i \overline{c_j} e^{i\langle t_i, x \rangle} \overline{e^{i\langle t_i, x \rangle}} \, \mu(dx) = \int \left| \sum_{i=1}^{n} c_i e^{i\langle t_i, x \rangle} \right|^2 \, \mu(dx) > 0$$

The following proposition requires the use of complex logarithms, which are not easily defined because the complex exponential function $e^z = e^{\Re z} e^{i\Im z} = e^{\Im z} (\cos(\Im z) + i\sin(\Im z))$ is not invertible as $e^z = e^{z+i2\pi k}$, $k \in \mathbb{Z}$.

9. **Proposition** (Cumulant function). Let $\phi \colon \mathbb{R}^n \to \mathbb{C}$ be the characteristic function of the probability measure μ and assume that $\phi(t) \neq 0$, $t \in \mathbb{R}^n$. There exists a unique continuous function $\psi \colon \mathbb{R}^n \to \mathbb{C}$ such that $\phi(t) = e^{\psi(t)}$ and $\kappa(0) = 0$. Such a function is Hermitian. It is called the cumulant function of μ .

Proof. In steps.

- (1) The cumulant function is unique. In fact, if $\phi(t) = e^{\psi_1(t)} = e^{\psi_2(t)}$, then $\Re \psi_i(t) = \log(|\phi(t)|)$, i = 1, 2, hence $e^{i\Im\psi_1(t)} = e^{i\Im\psi_2(t)}$, hence $\Im \psi_i(t) \Im \psi_2(t) = 2\pi k(t)$. The function $t \mapsto k(t)$ is continuous and integer valued on \mathbb{R}^n , then constant and equal 0 at t = 0.
- (2) The cumulant function is Hermitian. From $\phi(-t) = \overline{\phi(t)}$ it follows $e^{i\Im\psi(-t)} = e^{-i\Im\psi(t)}$, hence $\Im\psi(-t) + \Im\psi(t) = 2\pi k(t)$.

- (3) Continuous argument. As ϕ is never zero, we can define define $f(t) = \phi(t)/|\phi(t)|$ so that $f: \mathbb{R}^n \to \mathbb{T} = \{z \in \mathbb{C} | |z| = 1\}$ is continuous and f(0) = 1. A continuous argument of f is a continuous $\theta \colon \mathbb{R}^n \to \mathbb{R}$ such that $f(t) = e^{i\theta(t)}$ and $\theta(0) = 0$. If a continuous argument exist, then $\phi(t) = |\phi(t)| e^{i\theta(t)} = e^{\log(|\phi(t)|) + i\theta(t)} = e^{\psi(t)}$, with $\psi = \log |\phi| + i\theta$ continuous and $\psi(0) = \log 1 + \theta(0) = 0$.
- (4) Consider a continuous $f: B \to \mathbb{T}, B \subset \mathbb{R}^n$, such that $f(B) \neq \mathbb{T}$. Choose $e^{i\alpha} \in$ $\mathbb{T}\setminus f(B)$. Then the function $z\mapsto \arg(e^{-i\alpha}z)+\alpha$ is a continuous bijection of $\mathbb{T}\setminus\{e^{i\alpha}\}$ onto $\alpha, \alpha + 2\pi$, hence $f(t) = e^{i(\arg(e^{-i\alpha}f(t)) + \alpha)}$ so that $t \mapsto \theta(t) = \arg(e^{-i\alpha}f(t)) + \alpha$ is a continuous argument of f on B.
- (5) Let $f_1, f_2: B \to \mathbb{T}, B \subset \mathbb{R}^n$, be continuous. Assume f_1 has a continuous argument on B, $f_1 = e^{i\theta_1}$, and $f_1(t) + f_2(t) \neq 0, t \in B$. Then $f_1/f_2 \colon \mathbb{R}^n \to \mathbb{C}$ never equals -1, hence it has a continuous argument on B, $f_1/f_2 = e^{i\theta}$, so that $f_2 = e^{i(\theta_1 - \theta)}$.
- (6) Consider the function $f^R \colon B(R) = \{t | \|t\|^2 \leq R\} \times [0,1] \ni (t,\alpha) \mapsto f(\alpha t)$. The function f^R is continuous on a compact set, hence uniformly continuous, so that there exists a $n \in \mathbb{N}$ such that $\left| f^R(\alpha_1 t) - f^R(\alpha_2 t) \right| \leq 1$ for all t if $|\alpha_1 - \alpha_1| \leq 1/n$. Consider the sequence $f_j(t) = f^R(\frac{j}{n}t), j = 0, 1, ..., n$. We proceed by finite induction on j. if j = 0 then $f_0(t) = f(0) = 1$ and the continuous argument is $\theta_0 = 0$. If there exist a continuous argument on B(R) for f_i , j < n, as $|f_{j+1}(t) - f_j(t)| \leq 1$, the relation $f_{j+1}(t) + f_j(t) = 0$ is impossible, and the previous item shows that f_{j+1} has a continuous argument on B(R). Finally, note that $f_n = f^R$.
- (7) Each f has a continuous argument θ_R on B(R) and $\theta_{R_1}(t) = \theta_{R_2}(t)$ for all $t \in B(\min(R_1, R_2))$ because $B(\min(R_1, R_2))$ is connected and the uniqueness argument applies. Then the global continuous argument is defined by its restrictions.

10. Exercise.

(1) Let μ , ν , be probability measures on \mathbb{R}^n with characteristic function respectively $\check{\mu}, \check{\nu}.$ Then

$$\int e^{i\langle t,y\rangle}\check{\mu}(t) \ \nu(dt) = \int \check{\nu}(x+y) \ \mu(dx)$$

(2) Let $X \sim \mu$ independent of $Y \sim N(0,1)$ and $\nu \sim X + \sigma Y$. Let g_{σ} be the density of σY . Then $\nu = \mu * g_{\sigma}$ has density

$$p_{\sigma}(t) = \int g_{\sigma}(t-x) \ \mu(dx) = (2\pi)^{-n/2} \sigma^n \int \check{g}_{1/\sigma} \ \mu(dx) = \int e^{-i\langle x,y \rangle} \check{\mu}(x) g_{1/\sigma}(x) \ dx.$$

Proof.

- (1) We have $e^{i\langle t,y\rangle}\check{\mu}(t) = e^{i\langle t,y\rangle}\int e^{i\langle t,y\rangle}\mu(dy) = \int e^{i\langle t,x+y\rangle}\mu(dy)$ and we can take the integral with respect to ν to get $\int (e^{i\langle t,y \rangle} \check{\mu}(t)) \nu(dt) = \int \int e^{i\langle t,x+y \rangle} \mu(dy) \nu(dt) =$ $\int \check{\nu}(x+y) \ \mu(dx).$
- (2) From the properties of the Gaussian density and the previous equality.

11. **Proposition** (Inversion theorems). Let denote by μ be a probability measure on \mathbb{R}^n with inverse Fourier transform $\check{\mu}$.

(1) For all $f: \mathbb{R}^n \to \mathbb{R}$ continuous and with bounded support, $f \in C_{00}(\mathbb{R}^n)$,

$$\int f(x) \ \mu(dx) = (2\pi)^{-n} \lim_{T \to \infty} \int_{-T}^{T} \cdots \int_{-T}^{T} \left(\int f(s) \mathrm{e}^{-i\langle s,t \rangle} \ ds \right) \check{\mu}(t) \ dt.$$

- (2) The mapping $\mu \mapsto \check{\mu}$ is 1-to-1.
- (3) If $\check{\mu}$ is integrable, that is $\int |\check{\mu}(t)| dt < \infty$, then μ has a bounded and uniformly continuous density p with respect to the Lebesgue measure λ , and

$$p(x) = (2\pi)^{-n} \int \check{\mu}(t) e^{-i\langle i, x \rangle} dt = (2\pi)^{-n} \hat{\check{\mu}}(t).$$

Exercise.

(1) We have

$$\begin{split} \int_{-T}^{T} \cdots \int_{-T}^{T} \left(\int f(s) \mathrm{e}^{-i\langle s,t\rangle} \, ds \right) \check{\mu}(t) \, dt &= \\ \int_{-T}^{T} \cdots \int_{-T}^{T} \left(\int f(s) \mathrm{e}^{-i\langle s,t\rangle} \, ds \right) \left(\int \mathrm{e}^{i\langle t,y\rangle} \, \mu(dy) \right) \, dt &= \\ \int ds \, f(s) \int \mu(dy) \int_{-T}^{T} \cdots \int_{-T}^{T} dt \, \mathrm{e}^{i\langle t,y-s\rangle} &= \\ 2^n \int ds \, f(s) \int \mu(dy) \prod_{i=1}^n \frac{\sin(T(y_i - s_i))}{y_i - s_i} &= \\ 2^n \int \left(\int f(s) \prod_{i=1}^n \frac{\sin(T(y_i - s_i))}{y_i - s_i} \, ds \right) \, \mu(dy) = \quad T(y - s) = u \\ &= 2^n \int f(y - T^{-1}u) \prod_{i=1}^n \frac{\sin(u_i)}{u_i} \, du \to \pi^n \quad T \to \infty \end{split}$$

See another version in $[4, \S 16.6]$.

- (2) Follows from the previous inversion formula and the monotone class theorem. See also a direct proof based on Ex. . in [1, 14.1]
- (3) Use the approximation with the Gaussian kernel of Ex. ., see [2, Th 1.3.6]

- 12. **Definition** (Weak convergence, convergence in distribution).
 - (1) A sequence $(\mu_n)_{n \in \mathbb{N}}$ of probability measure on \mathbb{R}^n converges weakly to a probability measure μ if for all bounded and continuous $f : \mathbb{R}^n \to \mathbb{R}$ we have

$$\lim_{n \to \infty} \int f(x) \ \mu_n(dx) = \int f(x) \ \mu(dx), \quad f \in C_b(\mathbb{R}^n).$$

(2) If $(X_n)_{n\in\mathbb{N}}$ and X are random variables in \mathbb{R}^n we say that $\lim_{n\to\infty} X_n = X$ in distribution if $\lim_{n\to\infty} \mathbb{E}(f(X_n)) = \mathbb{E}(f(X)), f \in C_b(\mathbb{R}^n).$

13. **Proposition.** Let μ , ν , μ_n , $n \in \mathbb{N}$, be probability measures on \mathbb{R}^n . If $\lim_{n\to\infty} \mu_n = \mu$ weakly, then $\lim_{n\to\infty} \nu * \mu_n = \nu * \mu$ weakly.

Proof. If $f \in C_b(\mathbb{R}^n)$, then for all x we have $(y \mapsto f(x+y) \in C_b(\mathbb{R}^n)$, and dominated convergence implies that

$$\lim_{n \to \infty} \int f(z) \ (\nu * \mu_n)(dz) = \lim_{n \to \infty} \iint f(x+y) \ \nu(dx)\mu_n(dy) = \\ \lim_{n \to \infty} \int \left(\int f(x+y) \ \mu_n(dy) \right) \ \nu(dx) = \int \left(\lim_{n \to \infty} \int f(x+y) \ \mu_n(dy) \right) \ \nu(dx) = \\ \int \left(\int f(x+y) \ \mu(dy) \right) \ \nu(dx) = \int f(z) \ (\nu * \mu)(dz).$$

14. **Proposition** (Lévy continuity theorem). Let $(\phi_n)_{n\in\mathbb{N}}$ the the sequence of characteristic functions of the sequence of probability measures $(\mu_n)_{n\in\mathbb{Z}}$. If there exist the pointwise limit $\phi(t) = \lim_{n\to\infty} \phi_n(t), t \in \mathbb{R}^n$, and the limit function ϕ is continuous at 0, then ϕ is a characteristic function of a probability measure μ . In such a case, for all bounded continuous function $f : \mathbb{R}^n \to \mathbb{R}$, we have $\lim_{n\to\infty} \int f d\mu_n = \int f d\mu$, that is the sequence $(\mu_n)_{n\in\mathbb{N}}$ weakly converges to μ .

Proof. See [4, 18.1] or [1, Th. 19.1].

15. **Proposition.** The mapping $\mu \mapsto \check{\mu}$ is 1-to-1 from probability measure to positive definite functions whose value is 1 at 0.

Proof.

16. **Proposition** (Bochner theorem). If the function $\phi \colon \mathbb{R}^n \to \mathbb{C}$ is continuous, positive definite, and such that $\phi(0) = 1$, then there exists a probability measure μ such that $\phi = \check{\mu}$.

Exercise. Assume first that ϕ is integrable. Let g_{σ} be the density of the Gaussian N(0, $\sigma^2 I$) with characteristic function \check{g}_{σ} ,

$$g_{\sigma}(x) = (2\pi\sigma^2)^{-n/2} \mathrm{e}^{-\|x\|^2/2\sigma^2}, \quad \check{g}_{\sigma}(t) = \mathrm{e}^{-\sigma^2 \|t\|^2/2}.$$

Note that $\check{g}_{\sigma}(t) = (2\pi)^{-n/2} g_{1/sigma}(t)$. Let us compute

$$\phi * \check{g}_{\sigma}(t) = \int \phi(s) \check{g}_{\sigma}(t-s) \, ds$$
$$= \int \phi(s) \left(\int e^{i\langle t-s,x \rangle} g_{\sigma}(x) \, dx \right) \, ds$$
$$= \int e^{i\langle t,x \rangle} g_{\sigma}(x) \left(\int e^{-i\langle s,x \rangle} \phi(s) \right) \, dx$$
$$= \int e^{i\langle t,x \rangle} g_{\sigma}(x) \hat{\phi}(x) \, dx$$

As $\hat{\phi}$ is nonnegative being the Fourier transform of a positive definite function, and $\hat{\phi} \leq \int |\phi|$, we can renormalize $g_{\sigma} \cdot \hat{\phi}$ to get a probability density, so that

$$\frac{\phi * \check{g}_{\sigma}(t)}{\phi * \check{g}_{\sigma}(0)} = \frac{\phi * g_{1/\sigma(t)}}{\phi * g_{1/\sigma}(0)}$$

is a characteristic function. As $\sigma \to 0$, letting $\sigma s = u$,

$$\phi * \check{g}_{\sigma}(t) = \int \phi(t-s) \mathrm{e}^{-\sigma^2 \|s\|^2/2} ds = \sigma^{-1} \int \phi(t-\sigma^{-1}u) \mathrm{e}^{-\|u\|^2/2} du,$$
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$$\frac{\phi * \check{g}_{\sigma}(t)}{\phi * \check{g}_{\sigma}(0)} = \frac{\int \phi(t - \sigma^{-1}) \mathrm{e}^{-\|u\|^{2}/2} du}{\int \phi(-\sigma^{-1}u) \mathrm{e}^{-\|u\|^{2}/2} du} \to \frac{(2\pi)^{-n/2} f(t)}{(2\pi)^{-n/2} f(0)} = f(t).$$

If ϕ is not integrable, for each *n* the function $\phi_n : t \mapsto \phi(t)\check{g}_{1/n}(t)$ are positive definite, integrable, 1 at 0, and $\lim_{n\to\infty} \phi_n(t) = \phi(t)$.

17. **Definition** (Infinite divisibility).

- A random variable X is infinitely divisible if for all $n \in \mathbb{N}$ there exist IID random variables X_1, \ldots, X_n such that $X \sim X_1 + \cdots + X_n$.
- Equivalently, a probability measure μ is infinitely divisible if for all $n \in \mathbb{N}$ there exists a probability measure μ_n such that $\mu = (\mu_n)^{*n}$.
- Equivalently, a characteristic function ϕ is infinitely divisible if for all $n \in \mathbb{N}$ there exists a characteristic function ϕ_n such that $\phi = (\phi_n)^n$.

18. Proposition.

- (1) If the characteristic functions ϕ , ϕ_1 are infinitely divisible, then $\overline{\phi}$, $|\phi|^2$, $\phi\phi_1$ are infinitely divisible.
- (2) Each infinite divisible characteristic function ϕ has a cumulant function, $\phi = e^{\psi}$.

Exercise.

- (1) If ϕ and ϕ_1 are the characteristic functions of the random variables X and X_1 , respectively, then $\overline{\phi}$ is the characteristic function of -X, $|\phi|^2$ of X X', X' being an independent copy of X, $\phi\phi_1$ of $X + X_1$, X and X_1 independent.
- (2) Because of Proposition 9 we want $\phi(t) \neq 0, t \in \mathbb{R}^n$. For all $n \in \mathbb{N}$, let $\phi = (\phi_n)^n$. Then $|\phi|^2 = |\phi_n|^{2n}$, hence $|\phi|^{2/n} = |\phi_n|^2$ is a characteristic function for all n. The limit $\phi_*(t) = \lim_{n \to \infty} |f|^{2/n}$ is $\phi_*(t) = 1$ if $\phi(t) \neq 0$ and $\phi_*(t) = 0$ if $\phi(t) = 0$. As $\phi(t) = \neq 0$ in a neighborhood of 0, then ϕ_* is a characteristic function equal to 1 in a neighborhood of 0, hence it is a characteristic function, hence continuous, so that the case $\phi(t) = 0$ is impossible.

19. Exercise (Table of infinitely divisible characteristic functions).

Sampling. If ϕ_i , i = 1, ..., n, are infinitely divisible, then $\phi_1 \otimes \cdots \otimes \phi_n$ is infinitely divisible.

Affine transformation. If ϕ is infinitely divisible, then $s \mapsto e^{i\langle s, \mu \rangle} \phi(A^T s)$ is infinitely divisible. If ψ is the cumulant function, the transformed cumulant is $s \mapsto i \langle s, \mu \rangle + \psi(A^t s)$.

Dirac. The Dirac distribution δ_{μ} has characteristic function $\phi(t) = e^{i \langle t, \mu \rangle}$ and cumulant function $\psi(t) = i \langle t, \mu \rangle$.

Poisson. If $X \sim \text{Poisson}(\lambda)$, then X has characteristic function $\phi(t) = e^{\lambda(e^{it}-1)}$. It is infinitely divisible. The cumulant function is

$$\psi(t) = \lambda(e^{it} - 1) = \int (e^{ity} - 1) \ (\lambda\delta_1)(dy) = \int (e^{ity} - 1) \ \nu(dy)$$

Gaussian. If $X \sim \text{Normal}_n(0, I)$, then X has characteristic function $\phi(t) = e^{-\|t\|/2}$. It is infinitely divisible. The cumulant function is $\psi(t) = -\frac{1}{2} \|t\|$.

Gamma. If $X \sim \text{Gamma}(\gamma, \lambda)$, then X has characteristic function $\phi(t) = (1-i\lambda^{-1}t)^{-\gamma}$. It is infinitely divisible with *n*-th root $\phi_n(t) = (1-i\lambda^{-1}t)^{-\gamma/n}$ and $\psi(t) = \lim_{n \to \infty} n(\phi_n(t)-1)$. Let us consider the measure

$$\nu(dy) = (\gamma y)^{-1} \mathrm{e}^{-\lambda y} (y > 0) \, dy.$$

We have

$$\begin{split} \int (\mathrm{e}^{ity} - 1) \ \nu(dy) &= \gamma^{-1} \int_0^\infty \frac{\mathrm{e}^{ity} - 1}{y} \mathrm{e}^{-\lambda y} \ dy \\ &= i\gamma^{-1} \int_0^\infty \frac{\mathrm{e}^{ity} - 1}{iy} \mathrm{e}^{-\lambda y} \ dy \\ &= i\gamma^{-1} \int_0^\infty \left(\int_0^t \mathrm{e}^{isy} \ ds \right) \mathrm{e}^{-\lambda y} \ dy \\ &= i\gamma^{-1} \int_0^t ds \int_0^\infty dy \ \mathrm{e}^{-(\lambda - is)y} \\ &= i\gamma^{-1} \int_0^t (\lambda - is)^{-1} \ ds \\ &= -\gamma^{-1} \log \left(\lambda - is \right) \Big|_{s=0}^{s=t} \\ &= \log \left(\left(1 - i\lambda^{-1}t \right)^{-\gamma} \right). \end{split}$$

Hence

$$\psi(t) = \int (e^{ity} - 1) \ \nu(dy).$$

Note that the measure μ is infinite, but $y \mapsto e^{ity} - 1$ is integrable, because

$$\left|\mathbf{e}^{ity} - 1\right| = \left|i\int_{0}^{ty}\mathbf{e}^{iu} \, du\right| \le \left|ty\right|,$$

so that

$$\int \left| \mathrm{e}^{ity} - 1 \right| \ \nu(dy) \leqslant \gamma^{-1} \left| t \right| \int_0^\infty \mathrm{e}^{-\lambda y} \ dy = \left| t \right| / \gamma \lambda.$$

An other integral form is of interest. Let $h: \mathbb{R} \to \mathbb{R}$ be bounded, continuous, asymmetric h(-y) = -h(y), equal to the identity h(y) = y if |y| < 1 and constant for $|y| \ge 1$. Then

$$\int ith(y) \ \nu(dy) = it\gamma^{-1} \int_0^1 e^{-\lambda y} \ dy + it\gamma^{-1} \int_1^\infty y^{-1} e^{-\lambda y} \ dy = it\left(\lambda^{-1}(e^{-\lambda} - 1) + E_1(\lambda)\right) = it\mu$$

so that

$$\psi(t) = it\mu + \int (e^{ity} - 1 - ith(y)) \ \nu(dy).$$

The integrand in this form is a function that equals the Taylor remainder of order 2. In fact, if |y| < 1, then

$$\left| e^{ity} - 1 - ith(y) \right| = \left| e^{ity} - 1 - ity \right| \le \frac{t^2 y^2}{2} < \frac{t^2}{2}$$

With the same function as above,

$$\int ith(y) \ \nu(dy) = it\lambda \int_{|y|<1} y \ \sigma(dy) + it\lambda \int_{|y|\ge 1} \operatorname{sign}(y) \ \sigma(dy) = it\mu,$$

so that

$$\psi(t) = it\mu + \int (e^{ity} - 1 - ith(y)) \ \nu(dy).$$

Compound Poisson. Let N be a Poisson process with intensity λ and $(X_n)_{n \in \mathbb{N}}$ be IID with distribution σ . Assume N and $(X_n)_{n \in \mathbb{N}}$ independent. The process defined by $Y_t = \sum_{k=1}^{N_t} X_k$ is called *compound Poisson*. The characteristic function of Y_1 is

$$\phi(t) = \mathbf{E}\left(\mathbf{e}^{it\left(\sum_{k=1}^{N_1} X_k\right)}\right) = \sum_{n=0}^{\infty} \mathbf{E}\left(\mathbf{e}^{it\left(\sum_{k=1}^n X_k\right)} \left(N_1 = n\right)\right) = \sum_{n=0}^{\infty} (\check{\sigma}(t))^n \frac{\lambda^n}{n!} \mathbf{e}^{-\lambda} = \mathbf{e}^{\lambda((\check{\sigma}(t))^n - 1)}.$$

The cumulant function is

$$\psi(t) = \lambda((\check{\sigma}(t))^n - 1) = \int (\mathrm{e}^{ity} - 1) \ (\lambda\sigma)(dy) = \int (\mathrm{e}^{ity} - 1) \ \nu(dy).$$

Approximation by Compound Poisson. Let g_1 be the Gaussian density N(0,1). Then $\check{g}_1(x) = e^{-\|x\|^2/2}$ with cumulant function $\psi(x) = -\frac{1}{2} \|x\|^2$. Consider the infinite divisibility, $\frac{1}{m}\psi(x) = \frac{1}{2m} \|x\|^2$, with is the cumulant function of $e^{-\|x\|^2/2m} = \check{g}_{m^{-1/2}}(x)$, i.e. N(0, m^{-1}). Consider the distribution Compound Poisson CP($m, g_{m^{-1/2}}$), i.e. the distribution of $\sum_{k=1}^{N} X_i$, with $N \sim \text{Poi}(m)$, $(X_k)_k$ IID N(0, m^{-1}). The characteristic function is

$$t \mapsto \mathrm{e}^{m(\check{g}_{m^{-1/2}}(t)-1)} = \mathrm{e}^{m(\mathrm{e}^{-m^{-1}\|x\|^{2}/2}-1)},$$

and cumulant function

$$t \mapsto m(\mathrm{e}^{-m^{-1} \|x\|^2/2} - 1) = \int (\mathrm{e}^{i\langle t, x \rangle} - 1) \ (mg_{m^{-1/2}})(x) dx.$$

The left hand side converges to $-\|x\|^2/2$. Note the peculiar convergence of the right hand side, were $g_{m^{-1/2}}(x)dx$ converges weakly to δ_0 , while $mg_{m^{-1/2}}(x)dx$ does not converge weakly, as for all f integrable

$$\int f(x)mg_{m^{-1/2}}(x) \, dx = \int mf(m^{-1/2}x) \, g_1(x) \, dx$$

20. **Definition.** A function $f : \mathbb{R}^n \to \mathbb{C}$ is conditionally positive definite *if it is Hermitian* and $\sum_{i,j=1}^m c_i \overline{c_j} f(t_i - t_j) \ge 0$ for all $t_1, \ldots, t_m \in \mathbb{R}^n$, $c_1, \ldots, c_n \in \mathbb{C}$, $\sum_{i=1}^m c_i = 0$.

21. **Proposition** (Lévy-Kinchin formula). Let the μ be a probability measure on \mathbb{R}^n with characteristic function $\check{\mu}$ and cumulant function ψ , $\check{\mu}(t) = e^{\psi(t)}$. The following conditions are equivalent:

- (1) The probability measure is infinitely divisible.
- (2) For all $n \in \mathbb{N}$ the function $t \mapsto e^{\frac{1}{n}\psi(t)}$ is positive definite.
- (3) The cumulant function is conditionally positive definite.
- (4) The cumulant function has the following form

(1)
$$\psi(t) = i \langle t, \mu \rangle - \frac{1}{2} \langle \Gamma t, t \rangle + \int \left(e^{i \langle t, y \rangle} - 1 - i \langle t, h(y) \rangle \right) \nu(dy),$$

where

- (a) $h: \mathbb{R}^n \to \mathbb{R}^n$ is bounded, continuous, antisymmetric h(-t) = -h(t), equal to the identity h(t) = t in a neighborhood of 0^1 ;
- (b) $a, \mu \in \mathbb{R};$
- (c) Γ is a symmetric and positive definite $n \times n$ matrix;

¹Other choices are possible, cf [2], [3].

(d) ν is a positive measure on \mathbb{R}^n such that $\nu(\{0\}) = 0$, for all a > 0, $\nu\{t | ||t|| > a\} < 0$ ∞ , and $\int \|y\|^2 \nu(dy) < \infty$.

Given h, the decomposition is unique and it is called Lévy-Khinchin formula. The triple (μ, Γ, ν) is called the Lévy triple and ν is called the Lévy measure.

Elements of the proof.

 $(1 \Rightarrow 4)$ As ϕ is infinitely divisible, there exists characteristic functions ϕ_n such that $\phi(t) = (\phi_n(t))^n$, $n = 1, 2, \dots$ From the existence and uniqueness of the cumulant function, $\phi = e^{\psi}$, as $\phi(t) = (\phi_n(t))^n = (e^{\phi_n})^n = e^{n\psi_n(t)}$, we obtain $\psi_n = \frac{1}{n}\psi$. The function $t \mapsto e^{n(\phi_n(t)-1)}$ is the characteristic function of the compound Poisson CP (n, μ_n) , $\check{\mu}_n = \phi_n$, that is the characteristic function of $\sum_{k=0}^{N_n} X_k$, with $N_n \sim \text{Poi}(n), X_0 = 0, X_1, X_2, \dots$ IID $\mu_n, N_n, X_1, X_2, \dots$ independent. As

$$e^{n(\phi_n(t)-1)} = \exp\left(n(e^{\frac{1}{n}\psi(t)}-1)\right) = \exp\left(\frac{e^{\frac{1}{n}\psi(t)}-1}{1/n}\right) \to e^{\psi(t)} = \phi(t), \quad n \to \infty,$$

we have obtained that every infinitely divisible distribution is the limit of compound Poisson distribution, which are themselves infinitely divisible.

The cumulant function of the compound Poisson is

$$\psi_n^{(c)}(t) = \int (\mathrm{e}^{ity} - 1) \ \nu_n(dy) = i \langle t, \mu_n \rangle + \int (\mathrm{e}^{ity} - 1 - i \langle t, h(y) \rangle) \ \nu_n(dy),$$

with $\nu_n(dy) = n\mu_n(dy)$ and $\mu_n = \int h(y) \nu_n(dy)$. Note that the Gaussian term is missing. It is possibly produced in the limit as $n \to \infty$.

- $(4 \Rightarrow 3)$ The function $g(t) = i \langle t, \mu \rangle \frac{1}{2} \langle \Gamma t, t \rangle$ is such that for all $\alpha > 0$ the function $f(t) = e^{\alpha g(t)}$ is the characteristic function of N($\alpha \mu, \alpha \Gamma$), hence positive definite. It follows that $(e^{\alpha g(t)} - 1)/\alpha$ is conditionally positive definite and so is g(t) = $\lim_{\alpha \downarrow 0} (e^{\alpha g(t)} - 1)/\alpha$. A direct computation shows that the integral part of (1) is conditionally positive definite.
- $(3 \Rightarrow 2)$ The function $f(t) = \frac{1}{n}\psi(t)$ is conditionally positive definite for all n and f(0) = 0. Let us show that for all $t_1, \ldots, t_m \in \mathbb{R}$ the matrix $[f(t_i - t_j) - f(t_i) - \overline{f(t_j)}]_{i,i=1}^m =$ $[f(t_i-t_j)-f(t_i)-f(-t_j)]_{i,j=1}^m$ is positive definite. Chose $c_1,\ldots,c_m \in \mathbb{C}, t_{m+1}=0$, and c_{m+1} such that $\sum_{i=1}^{m+1} c_i = 0$. We have

$$\sum_{i,j=1}^{m} c_i \overline{c_j} (f(t_i - t_j) - f(t_i) - \overline{f(t_j)}) = \sum_{i,j=1}^{m} c_i \overline{c_j} f(t_i - t_j) - \sum_{i,j=1}^{m} c_i \overline{c_j} f(t_i) - \sum_{i,j=1}^{m} c_i \overline{c_j} f(-t_j) = \sum_{i,j=1}^{m} c_i \overline{c_j} f(t_i - t_j) + \overline{c_{m+1}} \sum_{i=1}^{m} c_i f(t_i) + c_{m+1} \sum_{j=1}^{m} \overline{c_j} f(-t_j) = \sum_{i,j=1}^{m+1} c_i \overline{c_j} f(t_i - t_j) \ge 0.$$

The matrix $[e^{a_{ij}}]$ is positive definite if the matrix $[a_{ij}]$ is positive definite, then the matrix $\left[e^{f(t_i-t_j)-f(t_i)-\overline{f(t_j)}}\right]_{i,j=1}^m$ s positive definite and also

$$\sum_{i,j=1}^{m} c_i \overline{c_j} e^{f(t_i - t_j)} = \sum_{i,j=1}^{m} (c_i e^{f(t_i)}) \overline{c_j e^{f(t_j)}} e^{f(t_i - t_j) - f(t_i) - \overline{f(t_j)}} \ge o$$
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which shows that the function $t \mapsto e^f(t) = e^{\frac{1}{n}\psi(t)}$ is positive definite. (2 \Rightarrow 1) The characteristic function $\phi(t) = e^{\psi(t)}$ is infinitely divisible because for all n $\phi(t) = (\mathrm{e}^{\frac{1}{n}\psi(t)})^n.$

Uniqueness Assume there are two representations, so that

$$\begin{split} i\langle t,\mu\rangle_1 &-\frac{1}{2}\langle \Gamma_1 t,t\rangle + \int \left(\mathrm{e}^{i\langle t,y\rangle} - 1 - i\langle t,h(y)\rangle\right) \ \nu_1(dy) = \\ &i\langle t,\mu\rangle_2 - \frac{1}{2}\langle \Gamma_2 t,t\rangle + \int \left(\mathrm{e}^{i\langle t,y\rangle} - 1 - i\langle t,h(y)\rangle\right) \ \nu_2(dy), \end{split}$$

hence

$$f(t) = \int \left(e^{i\langle t, y \rangle} - 1 - i \langle t, h(y) \rangle \right) \nu_1(dy) - \int \left(e^{i\langle t, y \rangle} - 1 - i \langle t, h(y) \rangle \right) \nu_2(dy),$$

where f(t) is a polynomial of degree at most 2. The convolution of this equality with $\chi = (\delta_0 - \delta_x)^{*3}$ gives

$$\int \hat{\chi}(y) \mathrm{e}^{i\langle x,y\rangle} \nu_1(dy) = \int \hat{\chi}(y) \mathrm{e}^{i\langle x,y\rangle} \nu_2(dy),$$

that is $(1 - e^{i\langle x, y \rangle})^3 \nu_1(dy) = (1 - e^{i\langle x, y \rangle})^3 \nu_2(dy)$, which implies $\nu_1 = \nu_2$. If the integrals are equal, then the first parts are equal.

References

- 1. Jean Jacod and Philip Protter, Probability essentials, second ed., Universitext, Springer-Verlag, Berlin, 2003. MR MR1956867 (2003m:60002)
- 2. Zoltán Sasvári, Multivariate characteristic and correlation functions, De Gruyter Studies in Mathematics, vol. 50, Walter de Gruyter & Co., Berlin, 2013. MR 3059796
- 3. Ken-iti Sato, Lévy processes and infinitely divisible distributions, Cambridge Studies in Advanced Mathematics, vol. 68, Cambridge University Press, Cambridge, 2013, Translated from the 1990 Japanese original, Revised edition of the 1999 English translation. MR 3185174
- 4. David Williams, Probability with martingales, Cambridge Mathematical Textbooks, Cambridge University Press, Cambridge, 1991.

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