

STOCHASTIC CALCULUS 2013
PART I

GIOVANNI PISTONE

1. WEAKLY ASSIGNMENT

Read [2, Ch 1-2]. All the random variables in the following are defined on the probability space $(\Omega, \mathcal{F}, \mathbb{P}(\cdot))$. Choose two among the exercises discussed below. Deadline: coming week.

2. EXPECTATION

Definition 1. If X is a nonnegative random variable, $X \in L^+$, its expectation is well defined and equal to

$$\mathbb{E}(X) = \int_0^{+\infty} \mathbb{P}(X > x) dx.$$

The mapping $x \mapsto \mathbb{P}(X > x)$ is nonincreasing, hence the integral is an ordinary integral. If $X \sim B(p)$, then $\mathbb{E}(X) = p$.

- Theorem 1.**
- (1) If $X \in L^+$, then $\mathbb{E}(X) = 0$ if, and only if $X = 0$ a.s.
 - (2) If $X, Y \in L^+$, then $X + Y \in L^+$ and $\mathbb{E}(X + Y) = \mathbb{E}(X) + \mathbb{E}(Y)$.
 - (3) If $X, Y \in L^+$ and $X \leq Y$, then $\mathbb{E}(X) \leq \mathbb{E}(Y)$, with equality if, and only if, $X = Y$ a.s.
 - (4) If $X_n \uparrow X$, $n \rightarrow \infty$, then $\mathbb{E}(X_n) \uparrow \mathbb{E}(X)$.

Proof. **Exercise 1** □

Definition 2. If X is a real random variable, and both $\mathbb{E}(X^+)$ and $\mathbb{E}(X^-)$ are finite, then $X \in L^1$ and its expected value is

$$\mathbb{E}(X) = \mathbb{E}(X^+) - \mathbb{E}(X^-).$$

Theorem 2. Let \mathcal{G} be a sub- σ -algebra of \mathcal{F} . For each $X \in L^1$ there exists a unique \tilde{X} such that

- (1) \tilde{X} is \mathcal{G} -measurable, and
- (2) $\mathbb{E}(YX) = \mathbb{E}(Y\tilde{X})$ if Y is bounded and \mathcal{G} -measurable.

The random variable \tilde{X} is called the conditional expectation of X given \mathcal{G} and it is denoted by $\mathbb{E}(X|\mathcal{G})$.

3. BREGMAN DIVERGENCE

Let f be (strictly) convex and C^1 on the open interval $I =]a, b[$, $a \in \{-\infty\} \cup \mathbb{R}$, $b > a$ and in $\mathbb{R} \cup \{+\infty\}$. Define the Bregman divergence between x and y in I as

$$d(y; x) = f(y) - f(x) - f'(x)(y - x).$$

The divergence is nonnegative and it is zero if, and only if, $x = y$. The mapping $y \mapsto d(y; x)$ is convex with derivative

$$\frac{d}{dy}d(y; x) = f'(y) - f'(x).$$

The mapping $x \mapsto d(y; x)$ is not convex in general. Check as **Exercise 2a** that

$$(1) \quad d(y; x) - d(y; z) = d(z; x) + (f'(z) - f'(x))(y - z)$$

Definition 3. If X and Y are random variables with values in I , the random variable $d(Y; X)$ is nonnegative and we can define their Bregman divergence by

$$D(Y; X) = \mathbb{E}(d(Y; X)).$$

The divergence is nonnegative, possibly $+\infty$, and it is zero if, and only if, $X = Y$ a.s.

If $f(x) = x^2$, then $I = \mathbb{R}$ and

$$\begin{aligned} d(y; x) &= y^2 - x^2 - 2x(y - x) \\ &= (y - x)^2 \end{aligned}$$

and

$$D(Y; X) = \mathbb{E}((Y - X)^2).$$

If $f(x) = x \ln x - x$ with $I =]0, +\infty[$, then

$$\begin{aligned} d(y; x) &= y \ln y - y - (x \ln x - x) - \ln x(y - x) \\ &= y(\ln y - \ln x) + x - y \\ &= y \ln \left(\frac{y}{x}\right) - y + x. \end{aligned}$$

If Q and P are positive densities, then

$$\begin{aligned} D(Q; P) &= \mathbb{E} \left(Q \ln \left(\frac{Q}{P} \right) - Q + P \right) \\ &= \mathbb{E} \left(Q \ln \left(\frac{Q}{P} \right) \right) - 1 + 1 \\ &= \mathbb{E} \left(Q \ln \left(\frac{Q}{P} \right) \right) = \mathbb{E}_Q \left[\ln \left(\frac{Q}{P} \right) \right]. \end{aligned}$$

If $P = 1$, then $D(Q; 1) = \mathbb{E}(Q \ln(Q))$ is called the entropy of Q . Check the keyword *minimal entropy measure* on Google, e.g. <http://www.math.ethz.ch/~mschweiz/Files/EMM-eqf.pdf>. For other examples of Bregman divergence, see http://en.wikipedia.org/wiki/Bregman_divergence.

Theorem 3. Let \mathcal{G} be a sub- σ -algebra of \mathcal{F} . Assume X and Y take values in I and X is \mathcal{G} -measurable. If $Y \in L^1$, the conditional expectation $\tilde{Y} = \mathbb{E}(Y|\mathcal{G})$ takes values in I and

$$D(Y; X) \geq D(Y; \tilde{Y}).$$

Viceversa, if \tilde{Y} is the minimum point of $X \mapsto D(Y; X)$ over all X taking values in I and \mathcal{G} -measurable, then $\hat{Y} = \mathbb{E}(Y|\mathcal{G})$.

Proof. **Exercise 2b**

- (1) If $I =]a, b[$ and a is finite, then $I \subset \{y: y - a > 0\}$. As $Y - a > 0$ a.s., it follows $\mathbb{E}(Y - a|\mathcal{G}) = \tilde{Y} - a > 0$ a.s.
- (2) From (1),

$$\begin{aligned} d(Y; X) - d(Y; \tilde{Y}) &= d(\tilde{Y}; X) + (f'(\tilde{Y}) - f'(X))(Y - \tilde{Y}) \\ &\geq (f'(\tilde{Y}) - f'(X))(Y - \tilde{Y}). \end{aligned}$$

If $(f'(\tilde{Y}) - f'(X))$ were bounded, the expected value of the inequality would be zero. If not, multiply the inequality by $(-n < f'(\tilde{Y}) - f'(X) < n)$.

- (3) If $D(Y; X) = d(Y; \tilde{Y})$ and $(f'(\tilde{Y}) - f'(X))$ were bounded, then $D(X; \tilde{Y}) = 0$. □

4. HERMITE POLYNOMIALS

Define

$$\begin{aligned} \delta f(x) &= x f(x) - f'(x) \\ &= -e^{x^2/2} \frac{d}{dx} \left(f(x) e^{-x^2/2} \right). \end{aligned}$$

If $Z \sim \mathcal{N}(0, 1)$, then

$$\mathbb{E}(g(Z)\delta f(Z)) = \mathbb{E}(dg(Z)f(Z)),$$

i.e. δ is the transpose of the derivative w.r.t. the standard Gaussian measure. Moreover, $d\delta - \delta d = \text{id}$.

Definition 4. The Hermite polynomials are

$$\begin{aligned} H_0 &= 1, \\ H_n(x) &= \delta^n 1, n > 0. \end{aligned}$$

The first Hermite polynomials are

$$\begin{aligned} H_1(x) &= x, \\ H_2(x) &= x^2 - 1, \\ H_3(x) &= x^3 - 3x, \\ H_4(x) &= x^4 - 6x^2 + 3, \dots \end{aligned}$$

Theorem 4. (1) Let $Z \sim N(0,1)$. The random variables $H_n(Z)$ are orthogonal, $E(H_n(Z)H_m(Z)) = 0$ if $n \neq m$, and $E(H_n(Z)^2) = n!$.
 (2) $dH_n = nH_{n-1}$,
 (3) $H_{n+1} = xH_n - nH_{n-1}$.

Proof. **Exercise 3.** There are various references, e.g. [1, 230–233] \square

Theorem 5. $(H_n(Z)/\sqrt{n!})$, $n = 0, 1, 2, \dots$, is an orthonormal basis of $L^2(\sigma(Z))$.

A proof is in the reference above. We skip it, but this example is of interest for us:

$$\exp\left(tZ - \frac{t^2}{2}\right) = \sum_{n=0}^{\infty} \frac{t^n}{n!} H_n(Z).$$

4.1. Gaussian quadrature.

Theorem 6. (1) Each Hermite polynomial H_n , $n \geq 1$, has n distinct real roots.
 (2) The roots of H_{n+1} are separated by the roots of H_n , $n \geq 1$.

Let f be a polynomial in one variable with real coefficients. By polynomial division

$$f(x) = q(x)H_n(x) + r(x)$$

where r has degree smaller than H_n and $r(x) = f(x)$ on $\{x: H_n(x) = 0\}$. The $n - 1$ degree polynomial r is the remainder.

It follows

$$\begin{aligned} E(f(Z)) &= E(q(Z)H_n(Z)) + E(r(Z)) \\ &= E(q(Z)\delta 1^n(Z)) + E(r(Z)) \\ &= E(d^n q(Z)) + E(r(Z)). \end{aligned}$$

Hence

$$E(f(Z)) = E(r(Z)) \quad \text{iff} \quad E(d^n q(Z)) = 0$$

Note that $d^n q(Z) = 0$ if and only if q has degree smaller than n and this is only if f has degree smaller or equal to $2n - 1$. The expected value can be zero in many other cases.

If r has degree less than n , then

$$r(x) = \sum_{i=1}^n r(x_i)L_i(x),$$

where x_1, \dots, x_n are the roots of $H_n(x) = 0$ and

$$L_i(x) = \frac{\prod_{j \neq i} (x - x_j)}{\prod_{j \neq i} (x_i - x_j)}$$

are the Lagrange polynomials. The quadrature formula follows:

$$\begin{aligned} E(f(Z)) &= E(r(Z)) \\ &= \sum_{i=1}^n r(x_i) E(L_i(Z)) \\ &= \sum_{i=1}^n r(x_i) w_{n,i}. \end{aligned}$$

The values of the weights $w_{n,i}$ are precomputed and available in numerical tables and numerical software. See e.g. the classical tables <http://people.math.sfu.ca/~cbm/aands/> and the relevant R functions.

Exercise 4 Write in detail the method of quadrature and compute numerically an example.

REFERENCES

1. Paul Malliavin, *Integration and probability*, Graduate Texts in Mathematics, vol. 157, Springer-Verlag, New York, 1995, With the collaboration of Hélène Airault, Leslie Kay and Gérard Letac, Edited and translated from the French by Kay, With a foreword by Mark Pinsky. MR MR1335234 (97f:28001a)
2. Steven E. Shreve, *Stochastic calculus for finance. II*, Springer Finance, Springer-Verlag, New York, 2004, Continuous-time models. MR 2057928 (2005c:91001)

COLLEGIO CARLO ALBERTO

E-mail address: giovanni.pistone@carloalberto.org